

Time reversal of a Markov Chain on trees in a population-genetical context

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1. The Evolving Moran Genealogy

2. EMG time reversal

3. Transient Markov Chains

The basic population model



Moran Model (P. Moran, 1958) in discrete time:

• Constant population size *n*, one split and one death per generation

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Moran Model (P. Moran, 1958) in discrete time:

- Constant population size *n*, one split and one death per generation
- The genealogy at any point $i > i^*$ can be interpreted as a Yule Tree T_i .



EMG definition and appearance

Definition

We call the process $(T_i)_{i \in \mathbb{N}}$, with $\Pr(T_{i+1} = T \in \mathcal{T}_n | T_i)$ dictated by the underlying Moran model, the *Evolving Moran Genealogy* (*EMG*).







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The derived process $(\mathcal{X}_i)_{i \in \mathbb{N}} \in \{0,1\}^{\mathbb{N}}$ with $\mathcal{X}_i = 1$ iff a root jump takes place from T_{i-1} to T_i is called *root jump process*.



Reversing the mechanics

Backward in time: "Merge" the most recent split and revive an individual by "regrafting" it into the tree at some branch segment *b*.







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With what probability to choose the regrafting segment?





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Kelly's Lemma

For a positive recurrent Markov Chain M on a state space S with transition matrix P, the matrix P' is the transition matrix of the reversed process iff

$$\forall \boldsymbol{s}, t \in \boldsymbol{S} : \boldsymbol{P}_{\boldsymbol{s}t}^{\flat} = \boldsymbol{P}_{\boldsymbol{t}\boldsymbol{s}} \cdot \frac{\pi_t}{\pi_{\boldsymbol{s}}}$$
(1)

where π is the stationary distribution of *M*.

(Kolmogorov 1936, Kelly 1979, Lovasz & Winkler 1983)



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Corollary (Loop Property)

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Closed loops have the same probability in forward and reversed process.



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The *EMG*[♭]

Lemma (Wiehe, W. 19)

The conditions of Kelly's Lemma are satisfied for the reversed process of the EMG iff the regrafting segment is chosen uniformly.



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Definition

We call the reversed process *EMG* backward in time, for short EMG^{\flat} .



Analyzing jump times via \textit{EMG}^{\flat}





Analyzing jump times via \textit{EMG}^{\flat}



Only one possibility of making the root jump in EMG^{\flat} ...



Analyzing jump times via EMG^b



Only one possibility of making the root jump in EMG^{\flat} ...

Under the EMG^{\flat} , the process of root jumps is geometric of intensity $2/n^2$.

Becomes a Poisson process of intensity 1 as $n \to \infty$ (Pfaffelhuber & Wakolbinger 2008).





Reversing a 2-allele Moran model

Assume one individual at time 0 has genotype *a*, all others have *b*.



Fixation curves backward in time



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Fixation curves backward in time

The 2-allele Moran model is transient.

Let *M* be a Markov chain on a discrete state space $S \cup \{a, b\}$ such that for each path ω :

$$\forall i \leq \alpha(\omega) : \omega_i = \mathbf{a}$$

 $\forall i \geq \beta(\omega) : \omega_i = \mathbf{b}$

where α,β are finite **a.s.**. Then

$$orall oldsymbol{s}, t \in oldsymbol{S}: oldsymbol{P}_{st}^{\flat} = oldsymbol{P}_{st} \cdot rac{\eta(t)}{\eta(s)}$$
 (2)

with $\eta(\boldsymbol{s})$ "average time" spent in state \boldsymbol{s}

$$\eta(\mathbf{k}) = \int_{\omega} \sum_{\alpha(\omega) < i < \beta(\omega)} \mathcal{X}_{\mathbf{s}}(\omega_i) \mathrm{d}\mathbf{p}(\omega)$$

(Hunt 1960)

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(Hunt 1960) "Loop Property" still true! Let *M* be a Markov chain on a discrete state space $S \cup \{a, b\}$ such that for each path ω :

$$\forall i \leq \alpha(\omega) : \omega_i = \mathbf{a}$$

 $\forall i \geq \beta(\omega) : \omega_i = \mathbf{b}$

$$P_{f(a)} := \begin{pmatrix} 0 & k/n & 1 \\ 1 & 0 & & & \\ 0 & \frac{k(n-k)}{n^2} & & \\ & \frac{k^2 + (n-k)^2}{n^2} & & \\ & \frac{k(n-k)}{n^2} & & \\ & & 0 & 1 \end{pmatrix} k/n$$

where α, β are finite **a.s.**. Then

$$\forall s, t \in S : P_{st}^{\flat} = P_{st} \cdot \frac{\eta(t)}{\eta(s)} \quad (2)$$
with $\eta(s)$ "average time" spent in state s

$$\eta(k) = \int_{\omega} \sum_{\alpha(\omega) < i < \beta(\omega)} \mathcal{X}_{s}(\omega_{i}) \mathrm{d}p(\omega)$$

$$P_{t(a)}^{\flat} := \begin{pmatrix} * & k/n & 1 \\ 0 & \frac{k(n-k+1)}{n^{2}} & \\ \frac{k(n-k+1)}{n^{2}} & \\ \frac{k(n-k+1)}{n^{2}} & \\ \frac{k(n-k+1)}{n^{2}} & \\ 0 & 0 & 0 \end{pmatrix} \stackrel{k/n}{}$$

(Hunt 1960) "Loop Property" still true!

Associated tree-valued processes



Backward process on Yule Trees for n = 4. Requires solving of a large nonlinear system (33 vars)



Birth-Death process with constant rate (MM1 queue) Assume

$$Pr(X_{i+1} = X_i + 1) = Pr(X_{i+1} = X_i - 1) = 1/2$$





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Choose an upper boundary n and condition the process on hitting n before 0.



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- Access to otherwise hidden statistical features

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Time reversal of Birth-Death processes

- Extendable to general birth and death probabilities and to continuous time
- Potential for application in phylogenetic/biostatistic simulation



Merci



Cher - If I could turn back time written by Diane Warren produced by Diane Warren and Guy Roche ©1989 Geffen Records

Lemma

In the time span between first occurrence and fixation of a mutant, the expected number of root jumps is $2 - \frac{2}{n}$.



Root jump distribution for n = 2 (blue), 5 (turquoise), 10 (green), 25 (red) and the